

Algebraic Aspects of Multiple Regression

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- 1 Introduction
- 2 Multiple Regression of Fuel Data
 - ANOVA for Model Comparison
 - Partial F -Tests: A General Approach
 - Testing Significance of a Single Term
 - Automatic Sequential Testing of Terms
- 3 Standard Errors for Coefficients
- 4 Standard Errors for Predicted and Fitted Values

Introduction

- In this module, we quickly review some fundamental aspects of the algebra of multiple regression.

Key Matrix Formulas

- We already saw in our treatment of the two-sample independent sample t -test how additional regressors can be tested for significance using the partial F -test for nested models, implemented in the R command `anova`.
- Now we present the formulas for the model, estimated coefficients, and standard errors.

Key Matrix Formulas

The Model and its Coefficients

- The multiple regression model can be written

$$E(Y|\mathbf{X}) = \mathbf{X}\boldsymbol{\beta} \quad (1)$$

$$\text{Var}(Y|\mathbf{X}) = \sigma^2 \quad (2)$$

- As in the simple regression model, the first column of \mathbf{X} is a column of 1's, and the first element of $\boldsymbol{\beta}$ is typically labeled β_0 , with subsequent elements labeled $\beta_1 \dots \beta_p$.
- Given a set of n criterion ("response") scores in \mathbf{y} and an $n \times p + 1$ set of predictor scores (including the intercept) in the matrix \mathbf{X} , the ordinary least squares estimates of $\boldsymbol{\beta}$ may be calculated as

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} \quad (3)$$

- The predicted scores are calculated as

$$\hat{\mathbf{y}} = \mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} = \mathbf{P}_x\mathbf{y} \quad (4)$$

- The residual scores are, of course,

$$\hat{\mathbf{e}} = \mathbf{y} - \hat{\mathbf{y}} = (\mathbf{I} - \mathbf{P}_x)\mathbf{y} = \mathbf{Q}_x\mathbf{y} \quad (5)$$

Multiple Regression of Fuel Data

Introduction

- Suppose we use $Dlic$, $Income$, $logMiles$, and Tax to predict $Fuel$.
- We begin by analyzing the scatterplot matrix.
- As we can see in the next slide, the potential predictors are only moderately related to $Fuel$.

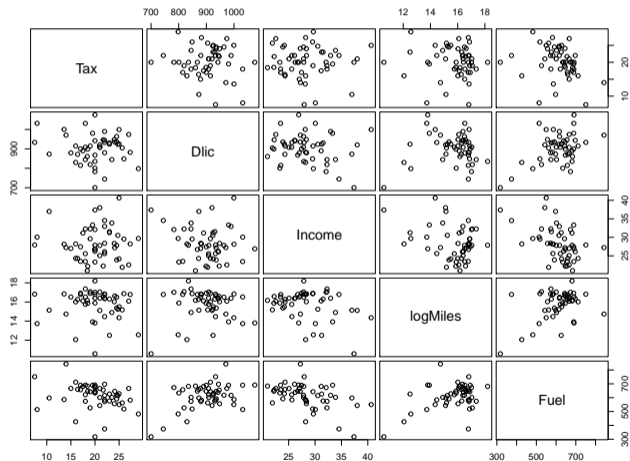
Multiple Regression of Fuel Data

Scatterplot Matrix Code

```
> data(fuel2001)
> fuel2001$Dlic <- 1000*fuel2001$Drivers/fuel2001$Pop
> fuel2001$Fuel <- 1000*fuel2001$FuelC/fuel2001$Pop
> fuel2001$Income <- fuel2001$Income/1000
> fuel2001$logMiles <- logb(fuel2001$Miles,2)
> f <- fuel2001[,c(7,9,3,10,9)]
> pairs(f,gap=0.4,cex.labels=1.5)
```

Multiple Regression of Fuel Data

Scatterplot Matrix



Multiple Regression of Fuel Data

Correlation Matrix

What we see in the scatterplot matrix is reflected in the matrix of intercorrelations.

```
> round(cor(f),4)
```

	Tax	Dlic	Income	logMiles	Fuel
Tax	1.0000	-0.0858	-0.0107	-0.0437	-0.2594
Dlic	-0.0858	1.0000	-0.1760	0.0306	0.4685
Income	-0.0107	-0.1760	1.0000	-0.2959	-0.4644
logMiles	-0.0437	0.0306	-0.2959	1.0000	0.4220
Fuel	-0.2594	0.4685	-0.4644	0.4220	1.0000

Multiple Regression of Fuel Data

Multiple Regression Output

- The next slide shows the output from the multiple regression for predicting *Fuel* from the 4 predictors.
- The far right column is the two-sided p -value for the t -statistic for each coefficient of a model term.
- In specifying the model, I use the specialized language used by R for setting up linear models. Each included term is assumed to have a coefficient, and the 1 explicitly indicates the intercept. R assumes an intercept is present. If you wish to specify a model with no intercept, you must include a -1 term.

Multiple Regression of Fuel Data

Multiple Regression Output

```
> attach(fuel2001)
> fuel.fit.all <- lm(Fuel~1 + Tax + Dlic + Income + logMiles)
> summary(fuel.fit.all)
```

Call:

```
lm(formula = Fuel ~ 1 + Tax + Dlic + Income + logMiles)
```

Residuals:

Min	1Q	Median	3Q	Max
-163.145	-33.039	5.895	31.989	183.499

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	154.1928	194.9062	0.791	0.432938
Tax	-4.2280	2.0301	-2.083	0.042873 *
Dlic	0.4719	0.1285	3.672	0.000626 ***
Income	-6.1353	2.1936	-2.797	0.007508 **
logMiles	18.5453	6.4722	2.865	0.006259 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 64.89 on 46 degrees of freedom

Multiple R-squared: 0.5105, Adjusted R-squared: 0.4679

F-statistic: 11.99 on 4 and 46 DF, p-value: 9.331e-07

Multiple Regression of Fuel Data

ANOVA for Model Comparison

- ANOVA is a key tool for comparing models.
- Define p' to be the number of terms in the regression model, including the intercept.
- As before, SSY is the sum of squared Y deviation scores, and RSS is the sum of squared residuals. Then

$$SS_{reg} = SSY - RSS \quad (6)$$

- To assess the overall significance of the prediction equation with 4 predictors, we follow the table shown below.

Source	df	SS	MS	F	p -value
Regression	p	SS_{reg}	$MS_{reg} = SS_{reg}/p$	$MS_{reg}/\hat{\sigma}^2$	
Residual	$n - p'$	RSS	$\hat{\sigma}^2 = RSS/(n - p')$		
Total	$n - 1$	SSY			

Multiple Regression of Fuel Data

ANOVA for Model Comparison

- The overall test for the *combined* significance of β_1 , β_2 , β_3 , and β_4 compares a model with only an intercept β_0 against a model with the intercept and all other terms.

```
> fuel.fit.intercept.only <- lm(Fuel~1)
> anova(fuel.fit.intercept.only,fuel.fit.all)
```

Analysis of Variance Table

Model 1: Fuel ~ 1

Model 2: Fuel ~ 1 + Tax + Dlic + Income + logMiles

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	50	395694				
2	46	193700	4	201994	11.992	9.331e-07 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Multiple Regression of Fuel Data

Partial F -Tests: A General Approach

- Actually, the F -tests we've been discussing so far are a special case of a general procedure for generating *partial F -tests* on a nested sequence of models.
- Suppose Model A includes Model B as a special case. That is, Model B is a special case of Model A where some terms have coefficients of zero. Then Model B is nested within Model A.
- If we define SS_a to be the sum of squared residuals for Model A, SS_b the sum of squared residuals for Model B, df_a to be $n - p_a$, where p_a is the number of terms in Model A including the intercept, and $df_b = n - p_b$, then to compare Model B against Model A, we compute the partial F -statistic as follows.

$$F_{df_b - df_a, df_a} = \frac{MS_{comparison}}{MS_{res}} = \frac{(SS_b - SS_a)/(p_a - p_b)}{SS_a/df_a} \quad (7)$$

Multiple Regression of Fuel Data

Testing Significance of a Single Term

- R does this model comparison for us using the `anova` function.
- Suppose we wish to test the significance of the *Tax* term when all the other 3 predictors are already in the model (along with the intercept).
- There are several ways we can do this in R.
- A direct way is to specify a second model without the *Tax* term and compare it to the model with the *Tax* term.

```
> Fuel.Fit.Without.Tax <- lm(Fuel ~ 1 + Dlic + Income + logMiles)
> Fuel.Fit.With.Tax <- lm(Fuel ~ 1 + Dlic + Income + logMiles + Tax)
> anova(Fuel.Fit.Without.Tax,Fuel.Fit.With.Tax)
```

Analysis of Variance Table

Model 1: Fuel ~ 1 + Dlic + Income + logMiles

Model 2: Fuel ~ 1 + Dlic + Income + logMiles + Tax

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	47	211964				
2	46	193700	1	18264	4.3373	0.04287 *

 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Multiple Regression of Fuel Data

Automatic Sequential Testing of Single Terms

- R will automatically perform a sequence of term-by-term tests on the terms in your model, *in the order they are listed in the model specification*.
- Just use the `anova` command on the single full model.
- You can prove for yourself (C.P.!) that the order of testing matters. The significance level for a term depends on the terms entered before it.

```
> anova(Fuel.Fit.With.Tax)
```

Analysis of Variance Table

Response: Fuel

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
Dlic	1	86854	86854	20.6262	4.019e-05	***
Income	1	59576	59576	14.1481	0.0004765	***
logMiles	1	37300	37300	8.8581	0.0046399	**
Tax	1	18264	18264	4.3373	0.0428733	*
Residuals	46	193700	4211			

 Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Standard Errors for Coefficients

- In a formula that is virtually identical in form to the simpler one for bivariate regression, the covariance matrix of the estimated regression coefficients is given by

$$\text{Var}(\hat{\beta}|\mathbf{X}) = \sigma^2(\mathbf{X}'\mathbf{X})^{-1} \quad (8)$$

- The unbiased estimate of σ^2 is

$$\hat{\sigma}^2 = \frac{\text{RSS}}{n - p'} = \frac{\text{RSS}}{n - (p + 1)} \quad (9)$$

- Consequently, the typical estimate for $\text{Var}(\hat{\beta}|\mathbf{X})$ is

$$\widehat{\text{Var}}(\hat{\beta}|\mathbf{X}) = \hat{\sigma}^2(\mathbf{X}'\mathbf{X})^{-1} \quad (10)$$

Standard Errors for Predicted and Fitted Values

Introduction

- You recall from our earlier discussion that there are two distinctly different standard errors that we can compute in connection with the regression line.
- One standard error, se_{pred} , deals with the situation where we have a new set of \mathbf{x} values, and we wish to compute the standard error for the value of \hat{y} computed from these values.
- Another standard error, se_{fit} , deals with the situation where we would like to compute a set of standard errors for the (population) fitted values on the regression line.

Standard Errors for Predicted and Fitted Values

Key Formulas

- Suppose we have observed, or will in the future observe, a new case with its own set of predictors that result in a vector of terms \mathbf{x}^* .
- We would like to predict the value of the response given \mathbf{x}^* .
- As in simple regression, the point prediction is $\tilde{y}^* = \mathbf{x}^{*'}\hat{\beta}$, and the standard error of prediction, $\text{sepred}(\tilde{y}^*|\mathbf{x}^*)$, is

$$\text{sepred}(\tilde{y}^*|\mathbf{x}^*) = \hat{\sigma} \sqrt{1 + \mathbf{x}^{*'}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}^*} \quad (11)$$

- Similarly, the estimated average of all possible units with a value \mathbf{x} for the terms is given by the estimated mean function at \mathbf{x} , $\hat{E}(Y|\mathbf{X} = \mathbf{x}) = \hat{y} = \mathbf{x}'\hat{\beta}$, with standard error given by

$$\text{sefit}(\hat{y}|\mathbf{x}) = \hat{\sigma} \sqrt{\mathbf{x}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}} \quad (12)$$

Standard Errors for Predicted and Fitted Values

Key Formulas

- A given software package may not produce all these estimates.
- If a program produces sefit but not sepred, the latter can be computed from the former from the result

$$\text{sepred}(\tilde{y}^* | \mathbf{x}^*) = \sqrt{\hat{\sigma}^2 + \text{sefit}(\tilde{y}^* | \mathbf{x}^*)^2} \quad (13)$$